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**Diagnostic value of conventional endoscopic ultrasound for lymph node metastasis in upper gastrointestinal neoplasia: A meta-analysis**

Chen C *et al*. EUS for upper gastrointestinal neoplasia LNM

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**Abstract**

BACKGROUND

Upper gastrointestinal neoplasia mainly includes esophageal cancer and gastric cancer, both of which have high morbidity and mortality. Lymph node metastasis (LNM), as the most common metastasis mode of both diseases, is an important factor affecting tumor stage, treatment strategy and clinical prognosis. As a new fusion technology, endoscopic ultrasound (EUS) is becoming increasingly used in the diagnosis and treatment of digestive system diseases, but its use in detecting LNM in clinical practice remains limited.

AIM

To evaluate the diagnostic value of conventional EUS for LNM in upper gastrointestinal neoplasia.

METHODS

Using the search mode of “MeSH + Entry Terms” and according to the predetermined inclusion and exclusion criteria, we conducted a comprehensive search and screening of the PubMed, Embase and Cochrane Library databases from January 1, 2000 to October 1, 2022. Study data were extracted according to the predetermined data extraction form. The quality of the included studies was assessed using the Quality Assessment of Diagnostic Accuracy Studies tool, and the results of the quality assessment were presented using Review Manager 5.3.5 software. Finally, Stata14.0 software was used for a series of statistical analyses.

RESULTS

A total of 22 studies were included in our study, including 2986 patients. The pooled sensitivity, specificity, positive likelihood ratio, negative likelihood ratio, diagnostic score and diagnostic odds ratio of conventional EUS in the diagnosis of upper gastrointestinal neoplasia LNM were 0.62 [95% confidence interval (CI): 0.50-0.73], 0.80 (95%CI: 0.73-0.86), 3.15 (95%CI: 2.46-4.03), 0.47 (95%CI: 0.36-0.61), 1.90 (95%CI: 1.51-2.29) and 6.67 (95%CI: 4.52-9.84), respectively. The area under the summary receiver operating characteristic curve was 0.80 (95%CI: 0.76-0.83). Sensitivity analysis indicated that the results of the meta-analysis were stable. There was considerable heterogeneity among the included studies, and the threshold effect was an important source of heterogeneity. Univariable meta-regression and subgroup analysis showed that tumor type, sample size and EUS diagnostic criteria were significant sources of heterogeneity in specificity (*P* < 0.05). No significant publication bias was found.

CONCLUSION

Conventional EUS has certain clinical value and can assist in the detection of LNM in upper gastrointestinal neoplasia, but it cannot be used as a confirmatory or exclusionary test.

**Key Words:** Endosonography; Esophageal neoplasms; Stomach neoplasms; Lymphatic metastasis; Diagnosis; Meta-analysis

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**Core Tip:** This meta-analysis examined the diagnostic value of conventional endoscopic ultrasound (EUS) for lymph node metastasis (LNM) in upper gastrointestinal neoplasia. The pooled analyses of 2986 patients from 22 studies performed herein show that conventional EUS has certain clinical value and can assist in the detection of LNM in upper gastrointestinal neoplasia, but it cannot be used as a confirmatory or exclusionary test. More high-quality studies are needed to further verify the diagnostic value of EUS and determine the best diagnostic criteria.

**INTRODUCTION**

Upper gastrointestinal neoplasia mainly includes esophageal cancer and gastric cancer, and their morbidity and mortality have long been among the top ten of the global cancer list, bringing great pain and burden to countries all over the world, and they are major global public health problems[1-4]. The onset of esophageal cancer and gastric cancer is hidden, and the best time for treatment has often been passed by the time they are clinically diagnosed. Lymph node metastasis (LNM), as the most common metastasis mode of both diseases, is an important basis for tumor staging, which largely determines the treatment plan and clinical prognosis of patients[5-8]. For patients with early tumor stages and no LNM, we can attempt endoscopic minimally invasive treatment, but for patients with LNM or advanced tumor stages, it is often necessary to consider comprehensive treatment, including radiotherapy, chemotherapy or surgery[9-11]. One study showed that when esophageal cancer has 0, 1-2 or more than 2 malignant lymph nodes, the median patient survival time is 66 mo, 14.5 mo or 6.5 mo, respectively[12]. Therefore, it is very important to accurately predict LNM.

Endoscopic ultrasound (EUS) combines the advantages of endoscopic technology and ultrasound technology; that is, it can evaluate the mucous membrane of the digestive tract with the naked eye, and it can also be used to detect the hierarchical structure and surrounding tissues of the digestive tract wall with ultrasound wave. EUS has the advantages of close observation distance, high resolution, low price and few adverse events. Since the 1980s, EUS has been gradually used in the diagnosis and treatment of many digestive system diseases, including the staging of gastrointestinal tumors, the identification of submucosal tumors, and the study of pancreatic or biliary tract diseases[13,14].Conventional EUS uses grayscale imaging technology for analysis, which can clearly display the status of lymph nodes near upper gastrointestinal neoplasia and identify the nature of lymph nodes according to the imaging features. When the endosonographic characteristics of lymph nodes are hypoechoic, round in shape, with a clear boundary and a size greater than 1 cm, the accuracy of conventional EUS in predicting malignant lymph nodes is more than 80%[15]. Some studies have shown that the ability of conventional EUS to detect LNM in upper gastrointestinal neoplasia is better than that of computed tomography and positron emission tomography, but some scholars believe that the diagnostic performance of conventional EUS is poor, and study results have varied widely[16-20]. The purpose of this meta-analysis was to explore the diagnostic value of conventional EUS for LNM in upper gastrointestinal neoplasia to guide clinical practice more effectively.

**MATERIALS AND METHODS**

This meta-analysis was conducted according to the Preferred Reporting Items for Systematic Review and Meta-Analysis statement[21,22]. The study protocol was registered in the PROSPERO database with the number CRD42022372170.

***Literature search***

We used the “MeSH + Entry Terms” search mode to conduct a comprehensive search of the PubMed, Embase and Cochrane Library databases before October 1, 2022. The specific search terms were as follows: (“esophageal neoplasms” OR “stomach neoplasms” OR “duodenal neoplasms”) AND (“lymphatic metastasis” OR “lymph nodes”) AND “endosonography” AND “diagnostic test search strategy”. We also manually searched the references of related studies.

***Study selection***

We imported all the retrieved articles into EndNote software (Version X9.1; Clarivate Analytics; Philadelphia, United States). Two researchers independently conducted study selection according to the predetermined inclusion and exclusion criteria with the process of identification, screening, eligibility and inclusion. To ensure consistency, we conducted exercises and tests before the formal selection, and the data were verified for internal consistency with the Kappa test during the selection process. If there was any disagreement, the decision was made by the two researchers together through consultation.

The study inclusion criteria were as follows: (1) Patients older than 18 who had recently been diagnosed with upper gastrointestinal neoplasia such as esophageal cancer, gastric cancer, and duodenal cancer; (2) LNM detected by conventional EUS; and (3) Diagnostic testing.

The exclusion criteria were as follows: (1) Studies published before 2000; (2) Case reports, conference abstracts, reviews, comments, letters, meta-analyses and systematic reviews; (3) Animal or *in vitro* models used as the objects of the study; (4) Sample size less than ten cases; (5) Inclusion of only stage cN0 patients; (6) Patients with other malignant tumors; (7) Patients who received or may have received preoperative neoadjuvant therapy; (8) Use of assistive technologies such as fine needle aspiration (FNA); (9) LNM diagnosis not made with postoperative pathological examination as the gold standard or radical surgery not performed for all patients; (10) Per patient not used as the analysis unit; (11) Inability to extract 2 × 2 tables of true positives (TP), false positives (FP), false negatives (FN), and true negatives (TN); (12) Repeated publication of the same data; and (13) Full text of English literature not found.

***Data extraction and quality assessment***

Two researchers independently extracted the study data using the predetermined data extraction form, and when they faced disagreement, a third researcher was consulted. Extracted data included: (1) Study characteristics such as first author, publication year, study country, study design and participating center; (2) Diagnostic test characteristics such as EUS model, EUS scan type, EUS examination method, EUS scan frequency, EUS diagnostic criteria, type and number of image interpretation experts, blinding, interval between EUS and surgery, gold standard and analysis unit; (3) Patient/tumor characteristics such as tumor type, tumor location, tumor stage, tumor histological type, neoadjuvant therapy, location of metastatic lymph nodes, age, sex and sample size; and (4) Statistical indicators such as TP, FP, FN, and TN. If the data were not reported directly, the sensitivity, specificity, accuracy and other indicators were used for reverse calculation.

Two researchers independently assessed the quality of the included studies using the Quality Assessment of Diagnostic Accuracy Studies (QUADAS-2) tool[23]. Disagreements were resolved through consultation. The results of the quality assessment were presented using Review Manager software (Version 5.3.5; Nordic Cochrane Centre; Copenhagen, Denmark).

***Statistical analysis***

All data evaluation and picture generation were completed by Stata software (Version 14.0; StataCorp LP; Texas, United States) using the MIDAS module of the bivariable mixed effects model. This model not only considers factors such as heterogeneity between studies, threshold effect and study size but also enables the bivariate nature of the original data to remain unchanged throughout the analysis process, thereby generating reliable statistical indicators. The pooled sensitivity, specificity, positive likelihood ratio (PLR), negative likelihood ratio (NLR), diagnostic score (DS) and diagnostic odds ratio (DOR) were calculated by drawing forest plots. The higher the values of DS and DOR were, the better the diagnostic effect of conventional EUS. The area under the curve (AUC) was obtained by drawing a summary receiver operating characteristic (SROC) curve, and the diagnostic performance was considered low, moderate, and high for AUCs of 0.5-0.7, 0.7-0.9 and 0.9-1.0, respectively. Fagan’s nomogram was used to reveal changes in the posttest probabilities. Likelihood ratio scatter diagram was used to evaluate the diagnostic performance of conventional EUS. Sensitivity analysis was used to assess the influence of individual studies on heterogeneity and observe the stability of the summary statistics. The threshold effect was determined according to whether the ROC plane showed a “shoulder-arm” point distribution. The *Q* statistical test was applied to assess the heterogeneity among the included studies, and heterogeneity was considered statistically significant when *P* < 0.05. The degree of heterogeneity was estimated based on the *I2* statistic, where *I2* < 25%, 25%-50%, 50%-75%, and ≥ 75% were considered low, moderate, substantial, and considerable heterogeneity, respectively. If the heterogeneity was high, meta-regression and subgroup analysis were used to explore the most significant source of heterogeneity. Publication bias was assessed with Deeks’ funnel plot, and *P* < 0.05 indicated statistical significance. The statistical methods of this study were reviewed by Professor Yao Zhang from the Department of Epidemiology, College of Preventive Medicine, Army Medical University of China.

**RESULTS**

***Study selection***

The study selection process is shown in Figure 1. A total of 729 articles were retrieved from three databases, and 22 articles were included in the manual search. The complete retrieval strategy of each database and manual search literature catalog can be found in Supplementary Table 1. Among them, 99 repeated articles were excluded after checking duplicates with EndNote software, 525 obviously irrelevant articles were excluded after reading the publication year, title and abstract, 8 articles were not published in English, 97 articles that did not meet the requirements were excluded after full-text reading, and 22 articles were included in the analysis according to the screening criteria[24-45]. In addition, the Kappa coefficient of the consistency test of the final selection results of the two researchers was 0.810 (*P* = 0.000).

***Characteristics and quality of the included studies***

This meta-analysis included 22 studies with 2986 patients. The basic information of the studies is shown in Table 1, and the detailed information is shown in Supplementary Table 2. Among them, the vast majority of studies were retrospective studies (21/22, 95.5%) and single center studies (20/22, 90.9%); ten studies were conducted in eastern countries, and twelve studies were conducted in western countries; the objects of twelve studies and ten studies were esophageal cancer and gastric cancer, respectively; none of the patients received neoadjuvant therapy before EUS and surgery, and the gold standard for the diagnosis of LNM in all studies was postoperative pathology. The results of the quality assessment based on the QUADAS-2 tool are shown in Figure 2, and detailed quality assessment information is shown in Supplementary Table 3.

***Meta-analysis outcomes***

**Primary outcomes:** The pooled sensitivity and specificity of conventional EUS in the diagnosis of upper gastrointestinal neoplasia LNM were 0.62 [95% confidence interval (CI): 0.50-0.73, *I2* = 91.50%] and 0.80 (95%CI: 0.73-0.86, *I2* = 86.10%), respectively, as shown in Figure 3A. According to the SROC curve, the AUC was 0.80 (95%CI: 0.76-0.83), as shown in Figure 4.

**Secondary outcomes:** The pooled PLR, NLR, DS, and DOR of conventional EUS in the diagnosis of upper gastrointestinal neoplasia LNM were 3.15 (95%CI: 2.46-4.03, *I2* = 61.17%), 0.47 (95%CI: 0.36-0.61, *I2* = 92.21%), 1.90 (95%CI: 1.51-2.29, *I2* = 60.94%) and 6.67 (95%CI: 4.52-9.84, *I2* = 99.99%), respectively, as shown in Figures 3B and C. The likelihood ratio scatter diagram showed that the summary PLR and NLR for the index test were in the fourth quadrant, suggesting that conventional EUS cannot be used as a confirmatory or exclusionary test, as shown in Figure 5. According to Fagan’s nomogram, when the EUS results were positive, the probability of diagnosing LNM increased from 50% to 76%; when the EUS results were negative, the probability of diagnosing LNM decreased from 50% to 32%, as shown in Figure 6.

***Validation of meta-analysis results***

**Sensitivity analysis:** We conducted sensitivity analysis by eliminating studies one by one, and the results showed that the pooled sensitivity change rate was ≤ 4.84% (*I2* change rate ≤ 2.75%), and the pooled specificity change rate was ≤ 2.50% (*I2* change rate ≤ 5.04%), indicating that the results of the meta-analysis were stable. Detailed data from the sensitivity analysis are shown in Supplementary Table 4.

**Heterogeneity:** Based on the *Q* statistical test and *I2* statistic, considerable heterogeneity was observed in the analysis for diagnostic sensitivity and specificity of conventional EUS. The ROC plane showed that the sensitivity was positively correlated with (1 - specificity), resulting in a “shoulder-arm” point distribution and indicating the existence of a threshold effect, as shown in Figure 7. According to the calculations from Stata software, the proportion of heterogeneity likely due to the threshold effect was 0.54. Because of the obvious heterogeneity among studies, we included five covariates: Tumor type (esophageal cancer or gastric cancer), study area (eastern country or western country), publication year (2010-2018 or 2000-2009), sample size (≥ 100 cases or < 100 cases) and EUS diagnostic criteria (criteria 1 or criteria 2). Univariable meta-regression and subgroup analysis were performed to identify the significant sources of heterogeneity. The results showed that tumor type, sample size and EUS diagnostic criteria were significant sources of heterogeneity in specificity (*P* < 0.05), as shown in Figure 8. The covariable assignment instructions are shown in Supplementary Table 5. The indicators for evaluating the diagnostic value of conventional EUS in each subgroup are shown in Table 2.

***Publication bias***

Deeks’ funnel plot showed that the distribution of all studies was relatively symmetrical; the asymmetry was not statistically significant (*P* = 0.654), indicating that there was no significant publication bias among the 22 studies, as shown in Figure 9.

**DISCUSSION**

In this meta-analysis, upper gastrointestinal neoplasia was considered as a whole, and the diagnostic value of conventional EUS for LNM was analysed. To ensure the reliability of the research results, we excluded studies that were old, had incomplete data, had small sample sizes and were published in languages other than English. The effects of incomplete surgical resection, neoadjuvant therapy, animal experiments, assistive technologies and other malignant tumors on the statistical results were excluded (which is also the reason for the small number of studies included in this meta-analysis). The results of conventional EUS were compared with postoperative pathology, and the data of 2986 patients in 22 studies were analyzed in detail.

The results of the quality assessment showed that many studies had a risk of bias, mainly because the proportion of retrospective studies was too high, and selective bias may have been present in the patient inclusion process. Four studies did not clearly describe the diagnostic criteria in the use of EUS, five studies unreasonably excluded some tumor patients (two studies limited the tumor location, two studies defined the location of metastatic lymph nodes, and one study only included esophageal cancer of ≤ pT2 stage), and fifteen studies did not specify the interval between EUS and surgery. However, we believe that since both esophageal cancer and gastric cancer are malignancies, examination and surgery should be arranged as soon as possible after clinical diagnosis. Although many studies did not specify the interval, it should not have had a significant impact on the research results. Regarding concerns regarding applicability, we think that the main reasons were the difference in diagnostic criteria of EUS and bias in patient selection. Therefore, caution should be taken in interpreting the results of the meta-analysis.

Due to the significant heterogeneity among the included studies, we used the bivariate mixed effect model to calculate statistics on various diagnostic evaluation indicators. The results showed that the pooled sensitivity and specificity of conventional EUS in diagnosing LNM in upper gastrointestinal neoplasia were 0.62 and 0.80, respectively, and the AUC of the SROC curve was 0.80, which indicated that the diagnostic value of conventional EUS was moderate. When EUS indicated positive or negative results, the posttest probability could be adjusted from the previous 50% to 76% and 32%, respectively. This result is meaningful for noninvasive examinations, indicating that conventional EUS has certain clinical value. However, it is undeniable that because the PLR < 10 and NLR > 0.1 in conventional EUS diagnosis and the DS and DOR were relatively small, this examination cannot be used to confirm or exclude LNM, which is consistent with the results of previous studies[46-49].It is not difficult to understand that, as with other imaging examinations, it is difficult for conventional EUS to reach such a high diagnostic level without obtaining lymph node tissue.

We explored the sources of heterogeneity among the included studies. First, we believe that the threshold effect could lead to heterogeneity because the 22 studies adopted a variety of EUS diagnostic criteria, and the “shoulder-arm” point distribution in the ROC plane also confirmed our view; the threshold effect might contribute 54% of the heterogeneity. Then, in view of the differences among the various studies, we included five covariables that could be easily grouped according to the collected data for meta-regression and subgroup analysis. Considering the limited number of studies, it would have been difficult to guarantee the accuracy of the statistical results of the simultaneous inclusion of five covariates, so we included individual covariates one by one for analysis. Although we were unable to identify significant sources of heterogeneity in sensitivity, we found that the significant sources of heterogeneity in specificity included tumor type, sample size, and EUS diagnostic criteria. However, after excluding the influence of the above factors, the heterogeneity within each subgroup was still obvious. Therefore, we have reason to believe that the heterogeneity was caused by a combination of factors. Many unincluded factors may also have been sources of heterogeneity; examples and the reasons they were not analyzed in detail included the study design, participating center and EUS scan type (because of the proportion imbalance within the group), the qualifications of the endoscopists, the EUS model and scan frequency (because of the complexity of the data), and the tumor stage, tumor location and location of metastatic lymph nodes (because these could not be accurately distinguished). We also found that the study area was not a significant source of heterogeneity, indicating that the diagnostic performance of conventional EUS for LNM in patients with upper gastrointestinal neoplasia in eastern and western countries is comparable. Publication year was also not a significant source of heterogeneity, indicating that the diagnostic performance of conventional EUS has not changed significantly in the past 20 years and that there may be technical barriers in conventional EUS that limit opportunities to significantly improve the ability of conventional EUS to identify malignant lymph nodes by relying solely on the diagnostic criteria of size, shape, boundary and echo.

Although the performance of conventional EUS in diagnosing LNM in upper gastrointestinal neoplasia remains nonideal, the diagnostic ability can be greatly improved with the assistance of EUS-guided FNA (EUS-FNA), EUS elastography (EUS-E) and contrast-enhanced EUS (CE-EUS)[50-52].EUS-FNA uses a slender biopsy needle to perform puncture biopsy for suspicious lesions under the guidance of EUS, which can provide histopathological information and is an accurate method to distinguish between benign and malignant lymph nodes. The sensitivity and accuracy of EUS-FNA in the diagnosis of regional LNM of upper gastrointestinal neoplasia are higher than those of conventional EUS. Chen *et al*[53] included 26 studies with 2753 patients for meta-analysis and found that the pooled sensitivity and specificity of EUS-FNA in differentiating benign and malignant lymph nodes were 87% and 100%, respectively, and the AUC was as high as 0.9912.EUS-E uses different colors to distinguish tissue hardness and displays different color images according to the elastic difference between lymph nodes and surrounding tissues, which can more clearly identify metastatic lymph nodes, improve the diagnostic performance of conventional EUS, and reduce unnecessary biopsies. Xu *et al*[54] included seven studies with 368 patients for meta-analysis and showed that the pooled sensitivity, specificity and AUC of EUS-E in the diagnosis of LNM were 88%, 85% and 0.9456, respectively.CE-EUS obtains enhanced images by using contrast agents, which can provide more information about the lesion tissue and can be used to identify metastatic lymph nodes. Lisotti *et al*[55] included four studies with 336 patients in their meta-analysis and indicated that the pooled sensitivity and specificity of CE-EUS in diagnosing LNM were 82.1% and 90.7%, respectively. However, our study only analyzed the diagnostic value of conventional EUS for LNM of upper gastrointestinal neoplasia, without considering the role of the above assistive technologies, which may underestimate the diagnostic value of EUS and affect the choice of clinicians. Therefore, we can carry out relevant studies in the next stage to evaluate the diagnostic value of various EUS assistive technologies in detail.

Our study only included patients who underwent radical surgery and did not receive preoperative neoadjuvant therapy, which inevitably led to case selection bias and excluded some patients with early tumors suitable for endoscopic treatment or patients with advanced tumors not suitable for surgical treatment. In addition, because preoperative neoadjuvant chemoradiotherapy can improve the treatment effect and prolong the survival time of some patients with upper gastrointestinal neoplasia, some patients with positive LNM may not have received the best treatment in this study. However, it is difficult to know the exact situation of LNM without obtaining complete pathological tissue, and preoperative neoadjuvant therapy will cause necrosis, fibrosis or inflammation of lymph nodes, which will affect the diagnostic effect of conventional EUS and the manifestations of postoperative histopathology. Therefore, to provide a reliable reference standard, we had to abandon the above cases in the study design stage.

Our study also has the following limitations. First, there were many retrospective studies with a long time span and use of different technologies and tools, which may have led to selection bias. Second, only studies published in English were included, which may have led to information bias. Third, there were differences in the study designs and implementation processes, which may have led to confounding bias. In addition, the significant heterogeneity may have affected the reliability and repeatability of the analysis results.

**CONCLUSION**

In conclusion, conventional EUS has certain clinical value and can assist in the detection of LNM in upper gastrointestinal neoplasia, but it cannot be used as a confirmatory or exclusionary test. There was great heterogeneity among the included studies, and more high-quality studies are needed to further verify the diagnostic value of EUS and determine its best diagnostic criteria. However, with the popularization of EUS technology, the use of assistive technologies such as EUS-FNA, EUS-E or CE-EUS, and the training of high-quality endoscopists, we believe that EUS will be increasingly valuable in the diagnosis of LNM in upper gastrointestinal neoplasia.

**ARTICLE HIGHLIGHTS**

***Research background***

Upper gastrointestinal neoplasia, mainly including esophageal cancer and gastric cancer, is a common cancer with high mortality. Accurate prediction of lymph node metastasis (LNM) is of great significance for guiding clinical treatment and improving the prognosis of patients. In recent years, endoscopic ultrasound (EUS) has become increasingly used in the diagnosis and treatment of gastrointestinal diseases, but its application in the detection of LNM remains limited.

***Research motivation***

Although previous studies have reported the diagnostic value of conventional EUS for LNM in upper gastrointestinal neoplasia, the relevant research conclusions were controversial, and the research results have varied widely. Therefore, we intend to further carry out this research through meta-analysis.

***Research objectives***

This study aimed to systematically search the literature and examine the diagnostic value of conventional EUS for LNM in upper gastrointestinal neoplasia by summarizing and analyzing the data.

***Research methods***

We conducted a comprehensive search and screening of the PubMed, Embase and Cochrane Library databases from January 1, 2000 to October 1, 2022. Then, relevant study data were extracted, and the quality of the included studies was assessed based on the Quality Assessment of Diagnostic Accuracy Studies tool. Afterward, a meta-analysis was performed using the statistical software Stata 14.0.

***Research results***

A total of 2986 patients in 22 studies were included. The results showed that the pooled sensitivity, specificity and area under the summary receiver operating characteristic curve of conventional EUS in the diagnosis of upper gastrointestinal neoplasia LNM were acceptable, which were 0.62 [95% confidence interval (CI): 0.50-0.73], 0.80 (95%CI: 0.73-0.86) and 0.80 (95%CI: 0.76-0.83), respectively. However, the pooled positive likelihood ratio and negative likelihood ratio were relatively poor, at 3.15 (95%CI: 2.46-4.03) and 0.47 (95%CI: 0.36-0.61), respectively. The pooled diagnostic score and diagnostic odds ratio were relatively small, at 1.90 (95%CI: 1.51-2.29) and 6.67 (95%CI: 4.52-9.84), respectively.

***Research conclusions***

Conventional EUS has certain clinical value and can assist in the detection of LNM in upper gastrointestinal neoplasia, but it cannot be used as a confirmatory or exclusionary test. More high-quality studies are needed to further verify the diagnostic value of EUS and determine the best diagnostic criteria.

***Research perspectives***

In the future, further clinical studies should be carried out to evaluate the diagnostic value of various EUS assistive technologies for LNM in upper gastrointestinal neoplasia and to evaluate the influence of neoadjuvant therapy on the diagnostic value of EUS for LNM in upper gastrointestinal neoplasia.

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**Footnotes**

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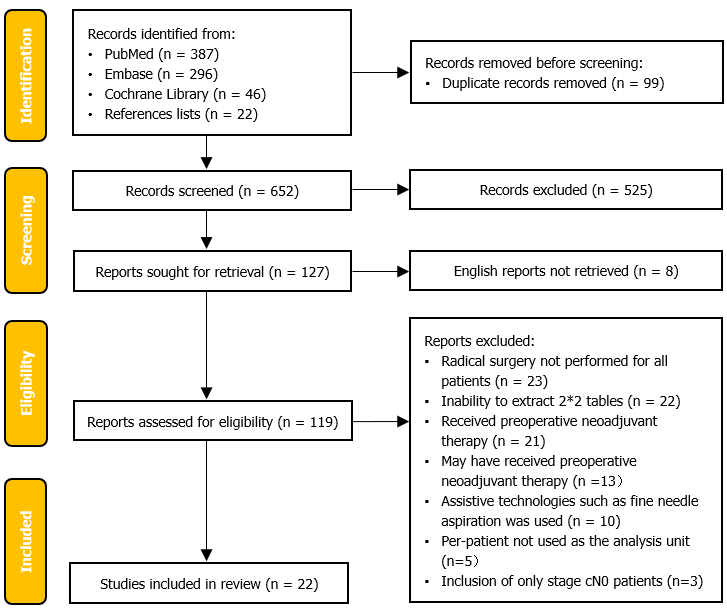
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Grade D (Fair): 0

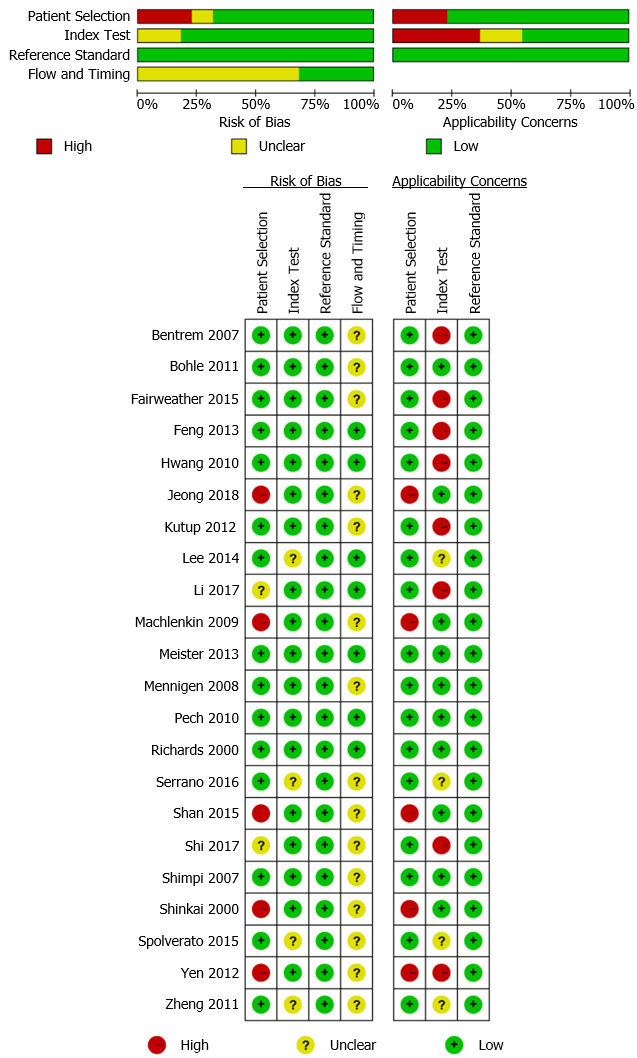
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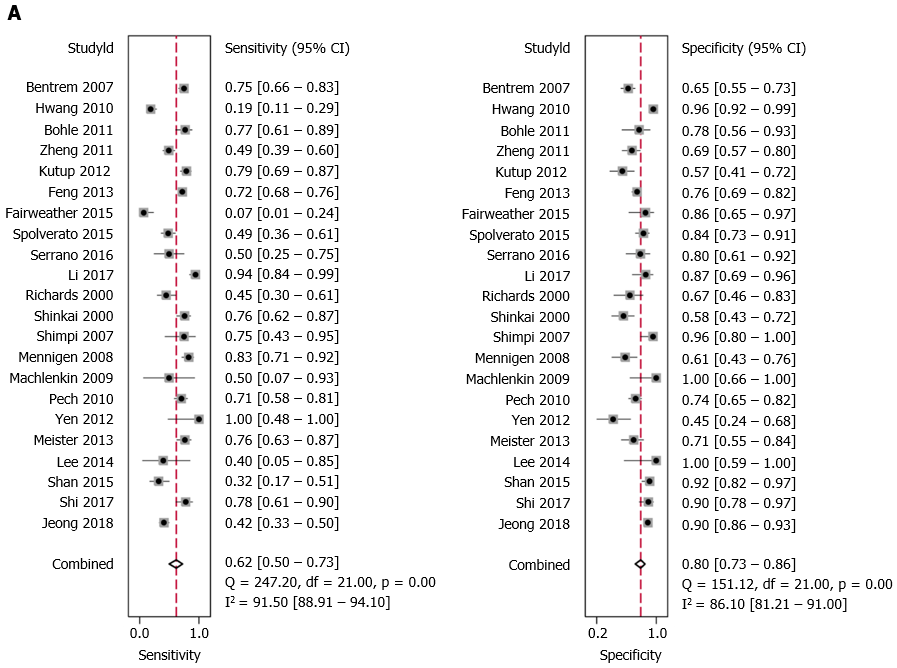
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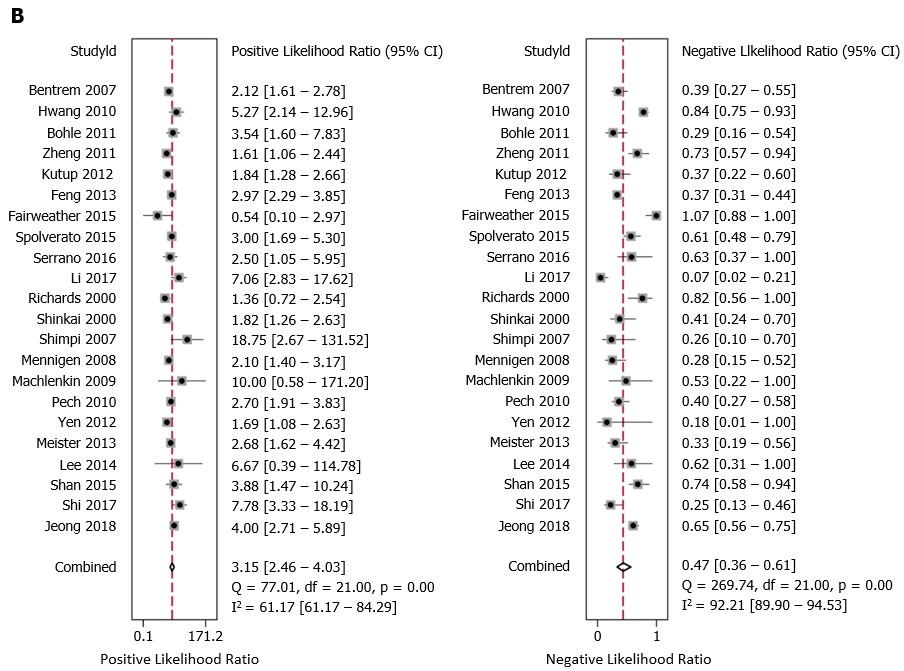


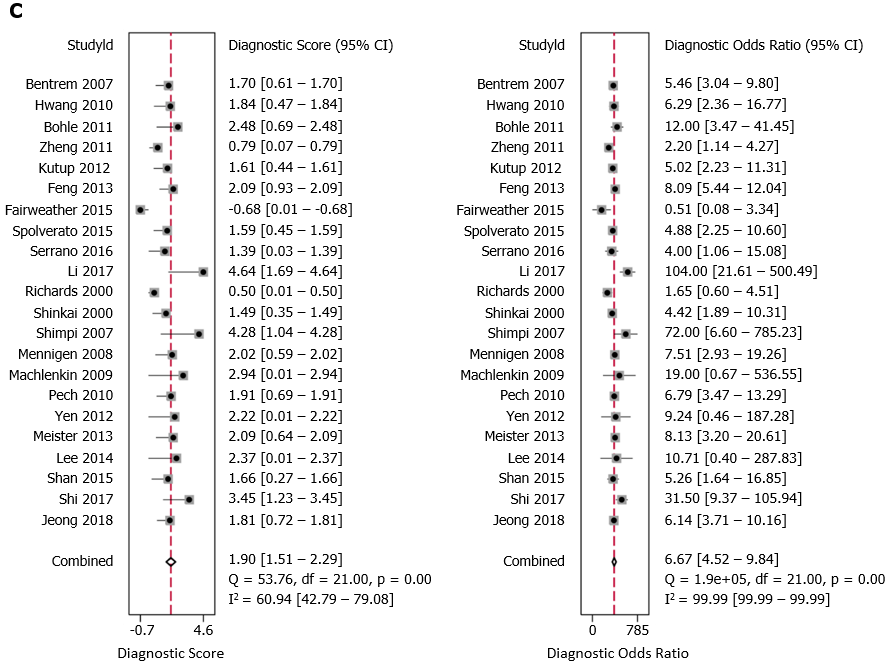
**Figure 1 Flow chart for study selection.**



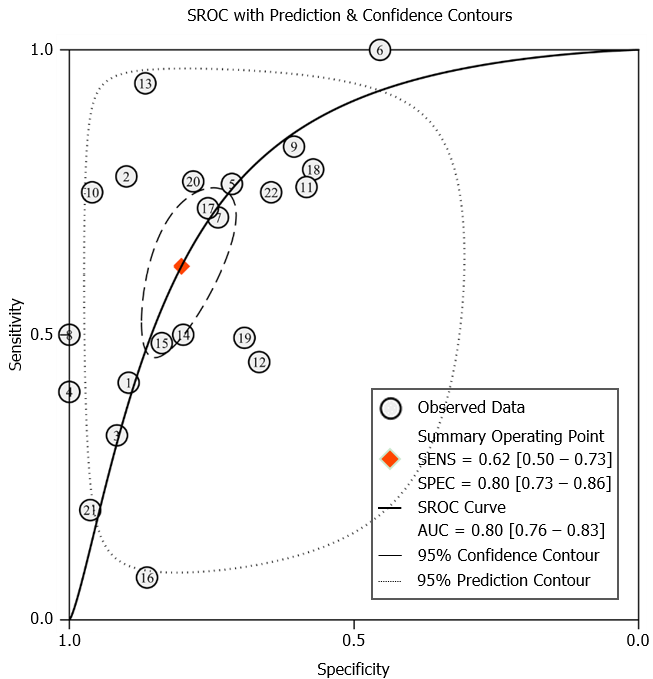
**Figure 2 Quality assessment of included studies based on the Quality Assessment of Diagnostic Accuracy Studies tool criteria.**



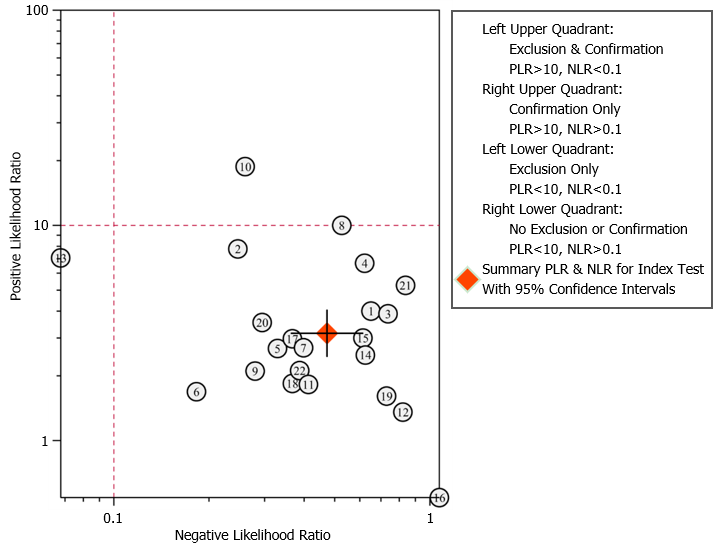




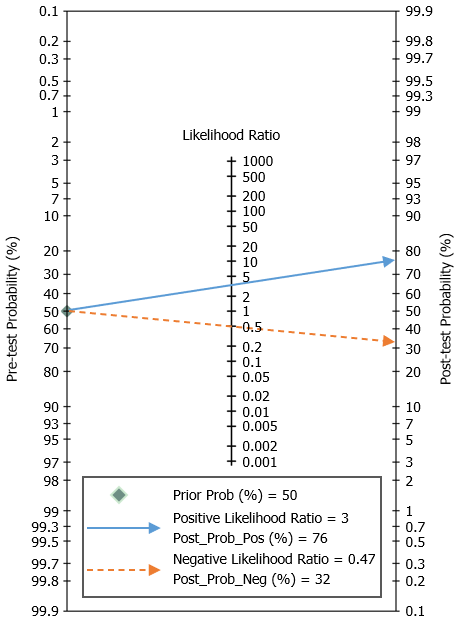
**Figure 3 Forest plots showing the pooled evaluation indicators and heterogeneity test results.** A: Pooled sensitivity and specificity; B: Pooled positive likelihood ratio and negative likelihood ratio; C: Pooled diagnostic score and diagnostic odds ratio. CI: Confidence interval.



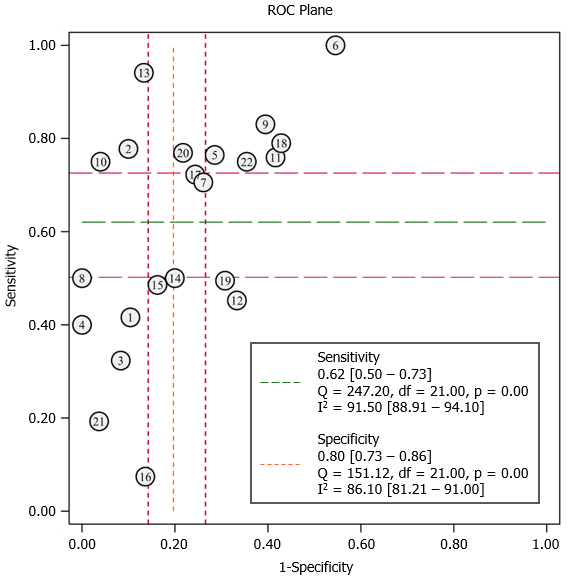
**Figure 4 Summary receiver operating characteristic curve for evaluating the diagnostic performance of conventional endoscopic ultrasound.** SROC: Summary receiver operating characteristic; SENS: Sensitivity; SPEC: Specificity; AUC: Area under the curve.



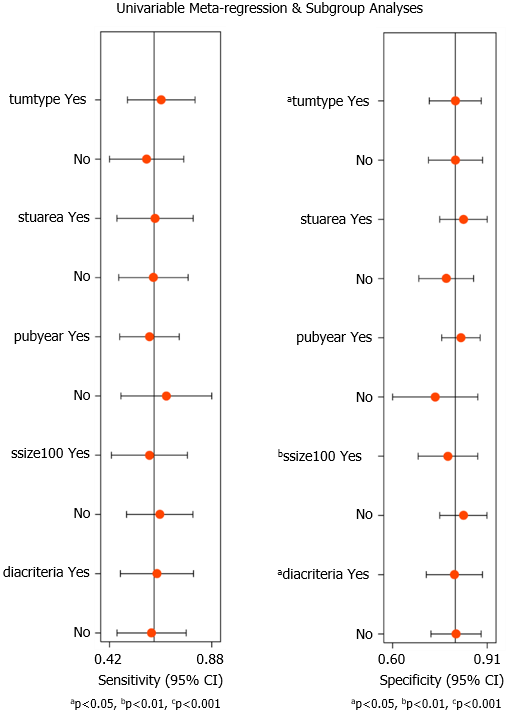
**Figure 5 Likelihood ratio scatter diagram for evaluating the diagnostic performance of conventional endoscopic ultrasound.** PLR: Positive likelihood ratio; NLR: Negative likelihood ratio.



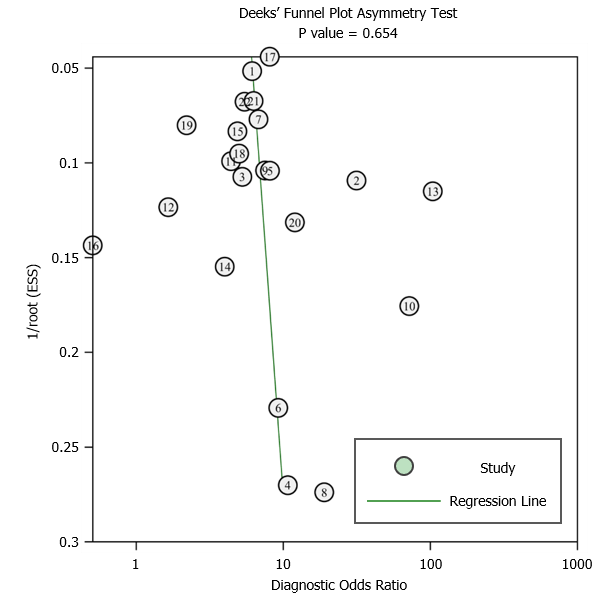
**Figure 6 Fagan’s nomogram for the diagnosis of lymph node metastasis with conventional endoscopic ultrasound.**



**Figure 7 Receiver operating characteristic plane for testing the threshold effect.** ROC: Receiver operating characteristic.



**Figure 8 Univariable meta-regression and subgroup analyses for finding sources of heterogeneity.** tumtype: Tumor type (Yes: Esophageal cancer; No: Gastric cancer); stuarea: Study area (Yes: Eastern countries; No: Western countries); pubyear: Publication year (Yes: 2010-2018; No: 2000-2009); ssize: Sample size (Yes: At least 100 cases; No: Less than 100 cases); diacriteria: Diagnostic criteria (Yes: Hypoechoic, round, well-defined margin, diameter ≥ 10mm; No: Others). CI: Confidence interval.



**Figure 9 Deeks’ funnel plot for assessing publication bias of the included studies.**

**Table 1 Characteristics of the included studies**

|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| **Ref.** | **Country** | **Study design** | **Center** | **EUS scan type** | **EUS scan frequency (MHz)** | **EUS diagnostic criteria1** | **Gold standard** | **Tumor type** | **Age2** **(yr)** | **Sample size (cases)** | **TP** | **FP** | **FN** | **TN** |
| Jeong *et al*[24], 2018 | Korea | Retrospective | 1 | Radial | 12/20 | Criteria 1 | Postoperative pathology | Esophageal cancer | 64 | 435 | 57 | 31 | 80 | 267 |
| Shi *et al*[25], 2017 | China | Retrospective | 1 | Radial | - | Criteria 2 | Postoperative pathology | Esophageal cancer | 59 | 86 | 28 | 5 | 8 | 45 |
| Shan *et al*[26], 2015 | China | Prospective | 1 | Radial | 7.5 | Criteria 1 | Postoperative pathology | Esophageal cancer | ≥ 44 | 94 | 11 | 5 | 23 | 55 |
| Lee *et al*[27], 2014 | Korea | Retrospective | 1 | Radial | 7.5/12/20 | Criteria 2 | Postoperative pathology | Esophageal cancer | 69 | 12 | 2 | 0 | 3 | 7 |
| Meister *et al*[28], 2013 | Germany | Retrospective | 5 | Radial | 20 | Criteria 1 | Postoperative pathology | Esophageal cancer | ≥ 34 | 93 | 39 | 12 | 12 | 30 |
| Yen *et al*[29], 2012 | China | Retrospective | 1 | Radial | 12/20 | Criteria 2 | Postoperative pathology | Esophageal cancer | ≥ 43 | 27 | 5 | 12 | 0 | 10 |
| Pech *et al*[30], 2010 | Germany | Retrospective | 1 | Radial | 7.5-10 | Criteria 1 | Postoperative pathology | Esophageal cancer | 64 | 179 | 48 | 29 | 20 | 82 |
| Machlenkin *et al*[31], 2009 | Israel | Retrospective | 1 | Radial | 7.5-12 | Criteria 1 | Postoperative pathology | Esophageal cancer | ≥ 28 | 13 | 2 | 0 | 2 | 9 |
| Mennigen *et al*[32], 2008 | Germany | Retrospective | 1 | - | 7.5/15 | Criteria 1 | Postoperative pathology | Esophageal cancer | 65 | 97 | 49 | 15 | 10 | 23 |
| Shimpi *et al*[33], 2007 | United States | Retrospective | 1 | Radial | 20 | Criteria 1 | Postoperative pathology | Esophageal cancer | - | 37 | 9 | 1 | 3 | 24 |
| Shinkai *et al*[34], 2000 | Japan | Retrospective | 1 | Radial | 7.5/12/15/20 | Criteria 1 | Postoperative pathology | Esophageal cancer | ≥ 42 | 102 | 41 | 20 | 13 | 28 |
| Richards *et al*[35], 2000 | United Kingdom | Retrospective | 1 | Radial | 7.5/12 | Criteria 1 | Postoperative pathology | Esophageal cancer | ≥ 35 | 69 | 19 | 9 | 23 | 18 |
| Li *et al*[36], 2017 | China | Retrospective | 1 | Radial | 5/7.5/10/12 | Criteria 2 | Postoperative pathology | Gastric cancer | 57 | 81 | 48 | 4 | 3 | 26 |
| Serrano *et al*[37], 2016 | United States | Retrospective | 1 | Radial | 7.5/10 | Criteria 2 | Postoperative pathology | Gastric cancer | ≥ 42 | 46 | 8 | 6 | 8 | 24 |
| Spolverato *et al*[38], 2015 | United States | Retrospective | 7 | - | - | Criteria 2 | Postoperative pathology | Gastric cancer | - | 144 | 34 | 12 | 36 | 62 |
| Fairweather *et al*[39], 2015 | United States | Retrospective | 1 | Radial | 5-10 | Criteria 2 | Postoperative pathology | Gastric cancer | 67 | 49 | 2 | 3 | 25 | 19 |
| Feng *et al*[40], 2013 | China | Retrospective | 1 | - | 5/7.5/12/15/20 | Criteria 2 | Postoperative pathology | Gastric cancer | 57 | 610 | 307 | 45 | 118 | 140 |
| Kutup *et al*[41], 2012 | Germany | Retrospective | 1 | Radial | 7.5/10/12 | Criteria 2 | Postoperative pathology | Gastric cancer | 61 | 123 | 64 | 18 | 17 | 24 |
| Zheng *et al*[42], 2011 | China | Retrospective | 1 | Radial | 7.5/12 | Criteria 2 | Postoperative pathology | Gastric cancer | 58 | 162 | 48 | 20 | 49 | 45 |
| Bohle *et al*[43], 2011 | Germany | Retrospective | 1 | Radial | 5-20 | Criteria 1 | Postoperative pathology | Gastric cancer | 63 | 62 | 30 | 5 | 9 | 18 |
| Hwang *et al*[44], 2010 | Korea | Retrospective | 1 | Radial | 5/7.5/12/20 | Criteria 2 | Postoperative pathology | Gastric cancer | ≥ 49 | 247 | 16 | 6 | 67 | 158 |
| Bentrem *et al*[45], 2007 | United States | Retrospective | 1 | - | 7.5-12 | Criteria 2 | Postoperative pathology | Gastric cancer | - | 218 | 81 | 39 | 27 | 71 |

1Criteria 1 (hypoechoic, round, well-defined margin, diameter ≥ 10 mm), Criteria 2 (others).

2Mean age or youngest age.

EUS: Endoscopic ultrasound; TP: True positives; FP: False positives; FN: False negatives; TN: True negatives.

**Table 2 Subgroup analysis of the diagnostic value of conventional endoscopic ultrasound for lymph node metastasis in upper gastrointestinal neoplasia**

|  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| **Subgroup** | | **Studies** | **Sensitivity (95%CI)** | **Specificity (95%CI)** | **PLR (95%CI)** | **NLR (95%CI)** | **DS (95%CI)** | **DOR (95%CI)** | **AUC (95%CI)** |
|  | All studies | 22 | 0.62 (0.50-0.73) | 0.80 (0.73-0.86) | 3.15 (2.46-4.03) | 0.47 (0.36-0.61) | 1.90 (1.51-2.29) | 6.67 (4.52-9.84) | 0.80 (0.76-0.83) |
| **Tumor type** | Esophageal cancer | 12 | 0.64 (0.51-0.76) | 0.81 (0.70-0.88) | 3.33 (2.27-4.87) | 0.44 (0.33-0.59) | 2.02 (1.53-2.50) | 7.52 (4.64-12.18) | 0.79 (0.75-0.83) |
| Gastric cancer | 10 | 0.59 (0.38-0.76) | 0.80 (0.71-0.87) | 2.95 (2.09-4.17) | 0.52 (0.34-0.79) | 1.74 (1.09-2.40) | 5.71 (2.96-10.99) | 0.79 (0.75-0.82) |
| **Study area** | Eastern country | 10 | 0.63 (0.42-0.80) | 0.84 (0.72-0.91) | 3.83 (2.46-5.95) | 0.44 (0.28-0.71) | 2.15 (1.48-2.83) | 8.62 (4.38-16.94) | 0.82 (0.79-0.85) |
| Western country | 12 | 0.61 (0.48-0.73) | 0.76 (0.69-0.82) | 2.57 (2.13-3.11) | 0.51 (0.38-0.67) | 1.62 (1.26-1.99) | 5.07 (3.53-7.30) | 0.77 (0.73-0.80) |
| **Publication year** | 2010-2018 | 16 | 0.60 (0.44-0.74) | 0.82 (0.75-0.88) | 3.36 (2.58-4.38) | 0.48 (0.35-0.68) | 1.94 (1.47-2.41) | 6.93 (4.34-11.08) | 0.81 (0.77-0.84) |
| 2000-2009 | 6 | 0.70 (0.56-0.81) | 0.75 (0.56-0.88) | 2.79 (1.50-5.18) | 0.41 (0.27-0.62) | 1.93 (1.01-2.85) | 6.87 (2.74-17.22) | 0.78 (0.74-0.81) |
| **Sample size** | ≥ 100 cases | 9 | 0.60 (0.46-0.73) | 0.78 (0.67-0.87) | 2.79 (2.14-3.65) | 0.51 (0.40-0.65) | 1.71 (1.45-1.97) | 5.52 (4.24-7.18) | 0.76 (0.72-0.80) |
| < 100 cases | 13 | 0.65 (0.61-0.90) | 0.83 (0.72-0.89) | 3.80 (2.45-5.89) | 0.42 (0.26-0.69) | 2.20 (1.42-2.99) | 9.04 (4.13-19.81) | 0.83 (0.80-0.86) |
| **EUS diagnostic criteria**1 | Criteria 1 | 10 | 0.62 (0.50-0.73) | 0.79 (0.71-0.86) | 3.04 (2.38-3.88) | 0.47 (0.37-0.60) | 1.86 (1.58-2.14) | 6.43 (4.84-8.54) | 0.78 (0.74-0.82) |
| Criteria 2 | 12 | 0.61 (0.41-0.79) | 0.81 (0.70-0.88) | 3.17 (2.19-4.60) | 0.48 (0.30-0.75) | 1.89 (1.22-2.57) | 6.63 (3.37-13.05) | 0.80 (0.76-0.83) |

1Criteria 1 (hypoechoic, round, well-defined margin, diameter ≥ 10mm), Criteria 2 (others).

PLR: Positive likelihood ratio; NLR: Negative likelihood ratio; DS: Diagnostic score; DOR: Diagnostic odds ratio; AUC: Area under the curve.